

# Transient random walk in $\mathbb{Z}^2$ with stationary orientations

Françoise Pène

Université de Bretagne Occidentale

UMR CNRS 6205

Département de Mathématiques, UFR Sciences et Techniques

6, avenue Victor Le Gorgeu, 29238 BREST Cedex 3, France

francoise.pene@univ-brest.fr

**Abstract.** *In this paper, we extend a result of Campanino and Petritis [1]. We study a random walk in  $\mathbb{Z}^2$  with a random environment. We suppose that the orientations of the horizontal floors are given by a stationary sequence of random variables  $(\xi_k)_{k \in \mathbb{Z}}$ . Once the environment fixed, the random walk can go either up or down or with respect to the orientation of the present floor (with the same probability). In [1], the  $(\xi_k)_{k \in \mathbb{Z}}$  is a sequence of independent identically distributed random variables. In [2], this result is extended to some cases of independent orientations chosen with stationary probabilities (not equal to 0 and to 1). In the present paper, we generalize the result of [1] to some cases when  $(\xi_k)_k$  is stationary. Moreover we give a slight extension of a result of [2].*

## 1 Introduction

We consider a random walk  $(M_n = (\tilde{X}_n, \tilde{Y}_n))_n$  in  $\mathbb{Z}^2$  with a random environment of a specific type. We suppose that  $M_0 = (0, 0)$ . Let  $(\xi_k)_{k \in \mathbb{Z}}$  be a stationary sequence of centered random variables with values in  $\{-1; 1\}$ . The orientations of the  $k^{th}$  horizontal floor of  $\mathbb{Z}^2$  is given by  $\xi_k$ . Once the environment fixed, the random walk  $(M_n = (\tilde{X}_n, \tilde{Y}_n))_n$  will be such that the distribution of  $M_{n+1} - M_n$  conditioned to  $\sigma(M_k; k = 0, \dots, n)$  is uniform on  $\{(0, 1); (0, -1); (\xi_{\tilde{Y}_n}, 0)\}$ .

In [1], Campanino and Petritis prove the transience of the random walk  $(M_n)_n$  when  $(\xi_k)_{k \in \mathbb{Z}}$  is sequence of independent identically distributed random variables.

In [2], the following situation is envisaged : Let  $(f_k)_{k \in \mathbb{Z}}$  be a stationary sequence of random variables with values in  $[0; 1]$  and with expectation equal to  $\frac{1}{2}$  defined on some probability space  $(M, \mathcal{F}, \nu)$ . Let us consider the probability space given by  $(\Omega_1 := M \times ]0; 1[^\mathbb{Z}, \mathcal{F}_2 := \mathcal{F} \otimes (\mathcal{B}(]0; 1[)^\otimes \mathbb{Z}, \nu_1 := \nu \otimes (\lambda)^\otimes \mathbb{Z})$ , where  $\lambda$  is the Lebesgue measure on  $]0; 1[$ . We define  $(\xi_{k, f_k})_{k \in \mathbb{Z}}$  on this space as follows :

$$\tilde{\xi}_{k, f_k}(\omega, (z_m)_{m \in \mathbb{Z}}) := 2 \cdot \mathbf{1}_{\{z_k \leq f_k(\omega)\}} - 1.$$

This means that, once a realization of  $(f_k)_k$  given, the horizontal floors are oriented independently; the  $k^{th}$  floor being oriented to the right with probability  $f_k$ . In [2], Guillotin and Le Ny prove that, if  $(\xi_k)_k = (\tilde{\xi}_{k, f_k})_k$ , then the corresponding random walk  $(M_n)_n$  is transient under the following condition :  $\int_M \frac{1}{\sqrt{f(1-f)}} d\nu < +\infty$ .

Let us notice that the  $(\xi_k)_k$  studied in [2] is stationary. Let us notice that, conversely, if  $(\xi_k)_k$  is stationary, then it can be described by the approach of [2] by taking  $f_k := \mathbf{1}_{\{\xi_k = 1\}} = \frac{1}{2}(\xi_k + 1)$ . But the method of [2] cannot be applied to a function  $f$  that can be equal to 0 or 1 with a non-null probability.

In this paper, we are interested in the case when  $(\xi_k)_{k \in \mathbb{Z}}$  is a stationary sequence of random variables satisfying some strong decorrelation properties.

We also end this paper with a short discussion about the model envisaged by Guillin and Le Ny. We prove that their result remains true if the condition  $\int_M \frac{1}{\sqrt{f(1-f)}} d\nu < +\infty$  is replaced by  $\int_M \frac{1}{[f(1-f)]^p} d\nu < +\infty$ , for some  $p > 0$ .

## 2 Result

**Theorem 1** *Let us suppose that :*

1. *we have :*

$$\sum_{p \geq 0} \sqrt{1+p} |\mathbb{E}[\xi_0 \xi_p]| < +\infty$$

$$\text{and } c'_0 := \sup_{N \geq 1} N^{-2} \sum_{k_1, k_2, k_3, k_4=0, \dots, N-1} |\mathbb{E}[\xi_{k_1} \xi_{k_2} \xi_{k_3} \xi_{k_4}]| < +\infty.$$

2. *There exists some  $C > 0$ , some  $(\varphi_{p,s})_{p,s \in \mathbb{N}}$  and some integer  $r \geq 1$  such that*

$$\forall (p,s) \in \mathbb{N}^2, \quad \varphi_{p+1,s} \leq \varphi_{p,s} \quad \text{and} \quad \lim_{s \rightarrow +\infty} s^6 \varphi_{r,s} = 0$$

*and such that, for all integers  $n_1, n_2, n_3, n_4$  with  $0 \leq n_1 \leq n_2 \leq n_3 \leq n_4$ , for all real numbers  $\alpha_{n_1}, \dots, \alpha_{n_2}$  and  $\beta_{n_3}, \dots, \beta_{n_4}$ , we have :*

$$\left| Cov \left( e^{i \sum_{k=n_1}^{n_2} \alpha_k \xi_k}, e^{i \sum_{k=n_3}^{n_4} \beta_k \xi_k} \right) \right| \leq C \left( 1 + \sum_{k=n_1}^{n_2} |\alpha_k| + \sum_{k=n_3}^{n_4} |\beta_k| \right) \varphi_{n_3-n_2, n_4-n_3}.$$

*Then the random walk  $(M_n)_n$  is transient.*

Let us notice that the hypotheses of this theorem are satisfied under the following  $\alpha$ -mixing condition :

$$\lim_{n \rightarrow +\infty} n^6 \sup_{p \geq 0, m \geq 0} \sup_{A \in \sigma(\xi_{-p}, \dots, \xi_0)} \sup_{B \in \sigma(\xi_n, \dots, \xi_{n+m})} |\mathbb{P}(A \cap B) - \mathbb{P}(A)\mathbb{P}(B)| = 0.$$

Moreover, in [4], we give examples of dynamical systems  $(M, \mathcal{F}, \nu, T)$  and of class of functions  $\mathcal{C}$  such that :

- If  $f : M \rightarrow [0; 1]$  belongs to the class  $\mathcal{C}$  and if  $\int_M f d\nu = \frac{1}{2}$  then  $(\xi_k := \tilde{\xi}_{k, f \circ T^k})_{k \in \mathbb{Z}}$  satisfies the hypotheses of our theorem 1.
- If  $g : M \rightarrow \{\pm 1\}$  belongs to the class  $\mathcal{C}$  and if  $\int_M g d\nu = 0$ , then  $(\xi_k := g \circ T^k)_{k \in \mathbb{Z}}$  satisfies the hypotheses of our theorem 1.

## 3 Proof of theorem 1

Let us define  $T_0 := 0$  and, for all  $n \geq 1$  :  $T_{n+1} := \inf\{k > T_n : \tilde{Y}_k \neq \tilde{Y}_{k-1}\}$ . According to lemma 2.5 of [1], we have the following result :

**Lemma 2** *If  $(M_{T_n})_{n \geq 0}$  is transient, then  $(M_n)_{n \geq 0}$  is transient*

Now, still following [1], we construct a realization of  $(M_{T_n})_n$  :

Let us consider a symmetric random walk  $(S_n)_n$  on  $\mathbb{Z}$  independent of  $(\xi_k)_{k \in \mathbb{Z}}$ . For any integer  $m \geq 1$  and any integer  $k$ , we define :

$$N_m(k) := \text{Card}\{j = 0, \dots, m : S_j = k\}.$$

Let us also consider a sequence of independent random variables  $(\zeta_i^{(y)})_{i \geq 1, y \in \mathbb{Z}}$  with geometric distribution with parameter  $\frac{1}{3}$ , and independent of  $(\xi_y)_{y \in \mathbb{Z}}$ .

**Lemma 3** *The process  $(X_n := \sum_{y \in \mathbb{Z}} \xi_y \sum_{i=1}^{N_{n-1}(y)} \xi_i^{(y)}, S_n)_{n \geq 1}$  has the same distribution as  $(M_{T_n})_n$ .*

In this lemma,  $\xi_i^{(y)}$  corresponds to the duration of the stay at the  $y^{th}$  horizontal floor during the  $i^{th}$  visit to this floor.

According to the Borel-Cantelli lemma, it suffices to prove that :  $\sum_{n \geq 1} \mathbb{P}(\{(X_n, S_n) = (0, 0)\}) < +\infty$ .

We follow the scheme of the proof of [1]. The difference will be in our way to estimate  $I_n^{(1)}$  and in the introduction of the sets  $U_n$ . We will consider  $\delta_1, \delta_2, \delta_3$ , and  $\gamma$  such that :

$0 < \delta_1 < 2\delta_2, \delta_1 + (\frac{27}{2} + 16)\delta_2 < \frac{1}{8}, \delta_3 > 0, \frac{1}{4} - 3\delta_2 < \delta_3 < \frac{1}{4} - \frac{5}{2}\delta_2 - \delta_1, \frac{\delta_3}{2} - 2\delta_2 < \beta < \frac{\delta_3}{2} - \delta_2, \max(\delta_1, \delta_2) < \gamma < \frac{1}{2} - 22 \max(\delta_1, \delta_2)$ . The idea is that  $\delta_1, \delta_2, \frac{1}{4} - \delta_3$  and  $\frac{1}{8} - \beta$  are positive numbers very close to zero.

As in [1, 2], let us define :

$$A_n := \{\omega \in \Omega : \max_{\ell \in \mathbb{Z}} N_{n-1}(\ell) \leq n^{\frac{1}{2} + \delta_2} \text{ and } \max_{k=0, \dots, n} |S_k| < n^{\frac{1}{2} + \delta_1}\}.$$

Moreover, we define :

$$U_n := \{\omega \in A_n : \forall x, y \in \mathbb{Z}, |N_{n-1}(x) - N_{n-1}(y)| \geq \sqrt{|x - y|n^{\frac{1}{2} + \gamma}}\}.$$

The sketch of the proof is the following :

1. As in proposition 4.1 of [1], we have :

$$\sum_{n \geq 1} \mathbb{P}(\{X_n = 0 \text{ and } S_n = 0\} \setminus A_n) < +\infty;$$

actually we have :  $\sum_{n \geq 1} \mathbb{P}(\{S_n = 0\} \setminus A_n) < +\infty$ ;

2. We will see in lemma 4 of the present paper that we have :

$$\sum_{n \geq 1} \mathbb{P}(A_n \setminus U_n) < +\infty.$$

Therefore, we have :

$$\sum_{n \geq 0} \mathbb{P}(\{X_n = 0 \text{ and } S_n = 0\} \setminus U_n) < +\infty;$$

3. Let us define  $B_n := \{\omega \in U_n : \left| \sum_{y \in \mathbb{Z}} \xi_y N_{n-1}(y) \right| > n^{\frac{1}{2} + \delta_3}\}$ . As in proposition 4.3 of [1], we have :

$$\sum_{y \in \mathbb{Z}} \mathbb{P}(B_n \cap \{X_n = 0 \text{ and } S_n = 0\}) < +\infty.$$

It remains to prove that :

$$\sum_{n \geq 0} \mathbb{P}(U_n \cap \{X_n = 0 \text{ and } S_n = 0\} \setminus B_n) < +\infty.$$

- (a) As in lemma 4.5 of [1], there exists a real number  $C > 0$  such that :

$$\sup_{\omega \in U_n \setminus B_n} \mathbb{P}(\{X_n = 0\} | (S_p)_{p \geq 1}, (\xi_k)_{k \in \mathbb{Z}}) \leq C \sqrt{\frac{\ln(n)}{n}}.$$

(b) We will prove that there exists some  $\tilde{\delta} > 0$  and some  $C' > 0$  such that :

$$\forall \omega \in U_n, \quad \mathbb{P}(U_n \setminus B_n | (S_p)_p)(\omega) \leq C' n^{-\tilde{\delta}}.$$

i. This probability is bounded by  $c' n^{\frac{1}{2} + \delta_3} I_n(\omega)$  with  $I_n(\omega) = I_n^{(1)}(\omega) + I_n^{(2)}(\omega)$  and

$$I_n^{(1)}(\omega) := \int_{\{|t| \leq n^{-\frac{1}{2} - \delta_3 + \delta_2}\}} \mathbb{E} \left[ e^{it \sum_{y \in \mathbb{Z}} \xi_y N_{n-1}(y)(\omega)} \middle| (S_p)_p \right] e^{-\frac{t^2 n^{1+2\delta_3}}{2}} dt$$

and

$$I_n^{(2)}(\omega) := \int_{\{|t| > n^{-\frac{1}{2} - \delta_3 + \delta_2}\}} \mathbb{E} \left[ e^{it \sum_{y \in \mathbb{Z}} \xi_y N_{n-1}(y)(\omega)} \middle| (S_p)_p \right] e^{-\frac{t^2 n^{1+2\delta_3}}{2}} dt.$$

ii. We will prove that  $n^{\frac{1}{2} + \delta_3} I_n^{(1)} = O(n^{-\delta})$  for some  $\delta > 0$  (see our lemma 5);

iii. On the other hand, following [1], we have :

$$n^{\frac{1}{2} + \delta_3} I_n^{(2)} \leq \int_{\{|s| > n^{\delta_2}\}} e^{-\frac{s^2}{2}} ds \leq 2n^{-\delta_2} e^{-\frac{n^{2\delta_2}}{2}}$$

(c) We have  $\mathbb{P}(Y_{n+1} = 0) \leq C'' n^{-\frac{1}{2}}$ .

(d) Hence we have :  $\mathbb{P}(U_n \cap \{X_n = 0 \text{ and } S_n = 0\} \setminus B_n) \leq C''' n^{-1-\delta} \sqrt{\ln(n)}$ .

We have to prove that the points 2 and 3(b)(ii) are true with our choices of parameters. All the other points are true for any positive  $\delta_1, \delta_2, \delta_3$  and for any sequence of random variables  $(\xi_k)_{k \in \mathbb{Z}}$  independent of  $(S_p)_p$ .

We notice that, for any integer  $n \geq 1$ , we have :  $\sum_{j=0}^{n-1} \xi_{S_j} = \sum_{k \in \mathbb{Z}} \xi_k N_{n-1}(k)$ . In our proof, we need some real numbers  $\delta_1, \delta_2, \delta_3, \delta_4, \beta, \gamma$  and  $\varepsilon > 0$ . We will suppose that :

$\delta_1 > 0, \delta_2 > 0, \delta_1 + (\frac{27}{2} + 16)\delta_2 < \frac{1}{8}, \delta_3 > 0, \delta_1 < \delta_4 < \frac{1}{4} - \delta_3 - \frac{5}{2}\delta_2, \frac{1}{4} - 3\delta_2 < \delta_3 < \frac{1}{4} - \frac{5}{2}\delta_2, \frac{5}{3}\delta_2 < \frac{1}{2}\delta_3, \frac{\delta_3}{2} - 2\delta_2 < \beta < \frac{\delta_3}{2} - \delta_2, \frac{5}{2}\delta_3 > \frac{1}{2} + 6\delta_2 + \delta_1, \max(\delta_1, \delta_2) < \gamma < \frac{1}{2} - 22 \max(\delta_1, \delta_2)$  and :

$$n^{\delta_1 + 11\delta_2} \sum_{m \geq \frac{(r+1)n^\beta}{2}} |\mathbb{E}[\xi_0 \xi_m]| = O(n^{-\varepsilon}).$$

(let us recall that we have supposed :  $\sum_{m \geq N} |\mathbb{E}[\xi_0 \xi_m]| \leq N^{-\frac{1}{2}} \sum_{m \geq N} \sqrt{m} |\mathbb{E}[\xi_0 \xi_m]|$ )

All this inequalities are true with the following choices of parameters :

$$\delta_1 = \frac{1}{3000}, \quad \delta_2 = \frac{1}{500}, \quad \delta_3 = \frac{1}{4} - \frac{11}{4}\delta_2, \quad \delta_4 = 1/2500, \quad \beta = \frac{\delta_3}{2} - \frac{3}{2}\delta_2 = \frac{1}{8} - \frac{23}{8}\delta_2, \quad \gamma = \frac{1}{4}.$$

**Lemma 4** *We have :*

$$\sum_{n \geq 1} \mathbb{P}(A_n \setminus U_n) < +\infty.$$

*Proof.* Let us consider any  $x, y \in \mathbb{Z}$  with  $x \neq y$  and  $|x - y| \leq 3n^{\frac{1}{2} + \delta_1}$ . For any integer  $j \geq 1$ , we define  $\tau_j(x)$  the time of the  $j^{\text{th}}$  visit of  $(S_p)_p$  to  $x$  and  $M_j(x, y)$  the number of visits of  $(S_p)_p$  to  $y$  between the times  $\tau_j(x)$  and  $\tau_{j+1}(x)$ . According to [3, 5], for any integer  $p \geq 1$ , there exists  $K_p > 0$  such that, for any  $x', y'$  we have :

$$\mathbb{E}[(M_j(x', y'))^p] \leq K_p |x' - y'|^{2p-1}.$$

According to [3], on the set  $\{\tau_1(x) \leq \tau_1(y)\}$ , we have :

$$(N_{n-1}(x) - N_{n-1}(y)) = \sum_{j=1}^{N_{n-1}(x)} (1 - M_j(x, y)) + \sum_{k=n}^{\tau_{N_{n-1}(x)+1}(x)} \mathbf{1}_{\{S_k=y\}}.$$

Let  $p$  be any positive integer. We have :

$$(N_{n-1}(x) - N_{n-1}(y))^{2p} \mathbf{1}_{\{\tau_1(x) \leq \tau_1(y)\}} \leq 2^{2p} \left[ \left( \sum_{j=1}^{N_{n-1}(x)} (1 - M_j(x, y)) \right)^{2p} + \left( \sum_{k=n}^{\tau_{N_{n-1}(x)+1}(x)} \mathbf{1}_{\{S_k=y\}} \right)^{2p} \right].$$

But, on  $A_n$ , since we have  $N_{n-1}(x) \leq n^{\frac{1}{2}+\delta_2}$ , we get :

$$\left( \sum_{k=n}^{\tau_{N_{n-1}(x)+1}(x)} \mathbf{1}_{\{S_k=y\}} \right)^{2p} \leq (M_{N_{n-1}(x)}(x, y))^{2p} \leq \sum_{j=1}^{\lfloor n^{\frac{1}{2}+\delta_2} \rfloor} (M_j(x, y))^{2p}.$$

Hence we have:

$$\begin{aligned} \mathbb{E} \left[ \left( \sum_{k=n}^{\tau_{N_{n-1}(x)+1}(x)} \mathbf{1}_{\{S_k=y\}} \right)^{2p} \mathbf{1}_{A_n} \right] &\leq n^{\frac{1}{2}+\delta_2} K_{2p} |x-y|^{2p-1} \\ &\leq n^{\frac{1}{2}+\delta_2} K_{2p} |x-y|^p \left( 3n^{\frac{1}{2}+\delta_1} \right)^{p-1} \\ &\leq K_{2p} 3^{p-1} |x-y|^p \left( n^{\frac{1}{2}+\max(\delta_1, \delta_2)} \right)^p. \end{aligned}$$

Moreover, on  $A_n$ , we have :

$$\left( \sum_{j=1}^{N_{n-1}(x)} (1 - M_j(x, y)) \right)^{2p} \leq \max_{k=1, \dots, \lfloor n^{\frac{1}{2}+\delta_2} \rfloor} \left( \sum_{j=1}^k (1 - M_j(x, y)) \right)^{2p}.$$

Since  $\left( \sum_{j=1}^k (1 - M_j(x, y)) \right)_{k \geq 1}$  is a martingale, according to a maximal inequality, we have :

$$\left\| \max_{k=1, \dots, \lfloor n^{\frac{1}{2}+\delta_2} \rfloor} \left( \sum_{j=1}^k (1 - M_j(x, y)) \right)^2 \right\|_{L^p} \leq \frac{p}{p-1} \max_{k=1, \dots, \lfloor n^{\frac{1}{2}+\delta_2} \rfloor} \left\| \left( \sum_{j=1}^k (1 - M_j(x, y)) \right)^2 \right\|_{L^p}.$$

Hence we have :

$$\mathbb{E} \left[ \left( \sum_{j=1}^{N_{n-1}(x)} (1 - M_j(x, y)) \right)^{2p} \mathbf{1}_{A_n} \right] \leq \left( \frac{p}{p-1} \right)^p \max_{k=1, \dots, \lfloor n^{\frac{1}{2}+\delta_2} \rfloor} \mathbb{E} \left[ \left( \sum_{j=1}^k (1 - M_j(x, y)) \right)^{2p} \right].$$

For any  $k = 1, \dots, \lfloor n^{\frac{1}{2}+\delta_2} \rfloor$ , we have :

$$\mathbb{E} \left[ \left( \sum_{j=1}^k (1 - M_j(x, y)) \right)^{2p} \right] = \sum_{l=1}^{2p} \sum_{\nu_1 + \dots + \nu_l = 2p; \min_i \nu_i \geq 1} \mathcal{M}_{\nu_1, \dots, \nu_l}^{2p} \sum_{j_1 < \dots < j_l} \prod_{m=1}^l \mathbb{E}[(1 - M_{j_m}(x, y))^{\nu_m}],$$

(since the  $M_{j_m}$  are independent) with  $\mathcal{M}_{\nu_1, \dots, \nu_l}^{2p} = \frac{(2p)!}{\prod_{i=1}^l \nu_i!}$ . Since  $\mathbb{E}[1 - M_j(x, y)] = 0$ , we have :

$$\mathbb{E} \left[ \left( \sum_{j=1}^k (1 - M_j(x, y)) \right)^{2p} \right] \leq$$

$$\begin{aligned}
&\leq \sum_{l=1}^{2p} \sum_{\nu_1+\dots+\nu_l=2p; \min_i \nu_i \geq 2} \mathcal{M}_{\nu_1, \dots, \nu_l}^{2p} \sum_{j_1 < \dots < j_l} \prod_{m=1}^l (2^{\nu_m} \mathbb{E}[1 + (M_{j_m}(x, y))^{\nu_m}]) \\
&\leq \sum_{l=1}^{2p} \sum_{\nu_1+\dots+\nu_l=2p; \min_i \nu_i \geq 2} \mathcal{M}_{\nu_1, \dots, \nu_l}^{2p} \sum_{j_1 < \dots < j_l} \prod_{m=1}^l (2^{\nu_m} (1 + K_{\nu_m} |x - y|^{\nu_m - 1})) \\
&\leq 2^{2p} \sum_{l=1}^{2p} |x - y|^{2p-l} (n^{\frac{1}{2} + \delta_2})^l \sum_{\nu_1+\dots+\nu_l=2p; \min_i \nu_i \geq 2} \mathcal{M}_{\nu_1, \dots, \nu_l}^{2p} \prod_{m=1}^l (1 + K_{\nu_m}) \\
&\leq \tilde{C}_p \sum_{l=1}^{2p} |x - y|^{2p-l} (n^{\frac{1}{2} + \delta_2})^l \\
&\leq \tilde{C}_p \sum_{l=1}^{2p} |x - y|^p (3n^{\frac{1}{2} + \delta_1})^{p-l} (n^{\frac{1}{2} + \delta_2})^l \\
&\leq 2p3^p \tilde{C}_p |x - y|^p (n^{\frac{1}{2} + \max(\delta_1, \delta_2)})^p.
\end{aligned}$$

Hence we get :

$$\mathbb{E}[(N_{n-1}(x) - N_{n-1}(y))^{2p} \mathbf{1}_{A_n}] \leq \tilde{C}_p' |x - y|^p (n^{\frac{1}{2} + \max(\delta_1, \delta_2)})^p.$$

Therefore, according to the Markov inequality, for any integer  $p \geq 1$ , we have :

$$\begin{aligned}
\mathbb{P}(A_n \setminus U_n) &\leq \sum_{x, y = -\lfloor n^{\frac{1}{2} + \delta_1} \rfloor}^{\lfloor n^{\frac{1}{2} + \delta_1} \rfloor} \mathbb{P}\left(A_n \cap \left\{|N_{n-1}(x) - N_{n-1}(y)| > \sqrt{|x - y| n^{\frac{1}{2} + \gamma}}\right\}\right) \\
&\leq \sum_{x, y = -\lfloor n^{\frac{1}{2} + \delta_1} \rfloor}^{\lfloor n^{\frac{1}{2} + \delta_1} \rfloor} \frac{\mathbb{E}[(N_{n-1}(x) - N_{n-1}(y))^{2p} \mathbf{1}_{A_n}]}{|x - y|^p (n^{\frac{1}{2} + \gamma})^p} \\
&\leq c_p \left(5n^{\frac{1}{2} + \delta_1}\right)^2 \left(n^{\max(\delta_1, \delta_2) - \gamma}\right)^p.
\end{aligned}$$

By taking  $p$  large enough, we get :  $\sum_{n \geq 1} \mathbb{P}(A_n \setminus U_n) < +\infty$ , *qed*.

### 3.1 Estimates on $U_n$

In this section, we suppose that we are in  $U_n$ . We will estimate :

$$I_n^{(1)}(\omega) := \int_{\{|t| \leq n^{-\frac{1}{2} - \delta_3 + \delta_2}\}} \left( \mathbb{E} \left[ e^{it \sum_{y \in \mathbb{Z}} \xi_y N_{n-1}(y)} \middle| (S_p)_p \right] (\omega) \right) e^{-\frac{t^2 n^{1+2\delta_3}}{2}} dt.$$

**Lemma 5** *There exists some real number  $\delta > 0$  such that we have :*

$$\sup_{n \geq 1} n^\delta \sup_{\omega \in U_n} n^{\frac{1}{2} + \delta_3} I_n^{(1)}(\omega) < +\infty.$$

We will use the following formula :

$$n^{\frac{1}{2} + \delta_3} I_n^{(1)}(\omega) = n^{\delta_2} \int_{\{|u| \leq 1\}} \left( \mathbb{E} \left[ e^{iun^{-\frac{1}{2} - \delta_3 + \delta_2} \sum_{y \in \mathbb{Z}} \xi_y N_{n-1}(y)} \middle| (S_p)_p \right] (\omega) \right) e^{-\frac{u^2 n^{2\delta_2}}{2}} du.$$

The main idea is to prove that, in  $I_n^{(1)}$ , we can replace the term :

$$B_n(u)(\omega) := \mathbb{E} \left[ e^{iun^{-\frac{1}{2} - \delta_3 + \delta_2} \sum_{y \in \mathbb{Z}} \xi_y N_{n-1}(y)} \middle| (S_p)_p \right] (\omega)$$

by

$$A_n(u)(\omega) := \exp \left( -\frac{u^2}{2n^{1+2\delta_3-2\delta_2}} \sum_{y,z} \mathbb{E}[\xi_y \xi_z] (N_{n-1}(y)(\omega))^2 \right).$$

More precisely we will prove the following :

**Lemma 6** *There exists a real number  $\delta_0 > 0$  such that we have :*

$$\sup_{n \geq 1} n^{\delta_0} \sup_{\omega \in U_n} n^{\delta_2} \int_{|u| \leq 1} |B_n(u)(\omega) - A_n(u)(\omega)| e^{-\frac{u^2 n^{2\delta_2}}{2}} du < +\infty. \quad (1)$$

Lemma 5 will be an easy consequence of lemma 6.

We will use the following notation :  $\sigma_\xi^2 := \sum_{m \in \mathbb{Z}} \mathbb{E}[\xi_0 \xi_m]$ .

### 3.1.1 Proof of lemma 6

We adapt the idea used in [4] to establish a result of convergence in distribution for  $\left( \sum_{k=0}^{n-1} \xi_{S_n} = \sum_y \xi_y N_{n-1}(y) \right)_{n \geq 1}$ .

Let  $n$  be an integer such that  $n^\beta \geq 2$ . Let us fix  $\omega \in U_n$  and  $u \in [-1; 1]$ . Let us recall that  $0 < \beta < \frac{\delta_3}{2} - \delta_2$

Let us define :

$$L_n := \left\lfloor \frac{2 \lfloor n^{\frac{1}{2} + \delta_1} \rfloor + 1}{\lfloor n^\beta \rfloor} \right\rfloor.$$

We notice that we have :  $L_n \leq 4n^{\frac{1}{2} + \delta_1 - \beta}$  since  $\lfloor n^\beta \rfloor \geq \frac{n^\beta}{2}$ .

$$\forall k = 0, \dots, L_n, \quad \alpha_{(k)} := - \lfloor n^{\frac{1}{2} + \delta_1} \rfloor + k \lfloor n^\beta \rfloor \text{ and } \alpha_{(L_n+1)} := \lfloor n^{\frac{1}{2} + \delta_1} \rfloor + 1;$$

$$b_k := \exp \left( i u n^{-\frac{1}{2} - \delta_3 + \delta_2} \sum_{y=\alpha_{(k)}}^{\alpha_{(k+1)}-1} \xi_y N_{n-1}(y) \right);$$

$$a_k := \exp \left( -\frac{u^2}{2n^{1+2\delta_3-2\delta_2}} \sum_{y=\alpha_{(k)}}^{\alpha_{(k+1)}-1} \sigma_\xi^2 (N_{n-1}(y))^2 \right).$$

We have to estimate :

$$n^{\delta_2} \left| \mathbb{E} \left[ \prod_{k=0}^{L_n} b_k \middle| (S_p)_p \right] (\omega) - \prod_{k=0}^{L_n} a_k(\omega) \right|.$$

Hence it is enough to estimate :

$$n^{\delta_2} \sum_{k=0}^{L_n} \left| \mathbb{E} \left[ \left( \prod_{m=0}^{k-1} b_m \right) (b_k - a_k) \left( \prod_{m'=k+1}^{L_n} a_{m'} \right) \middle| (S_p)_p \right] (\omega) \right|.$$

- We explain how we can restrict our study to the sum over the  $k$  such that  $(r+1)^4 \leq k \leq L_n - 1$ . Indeed, the number of  $k$  that does not satisfy this is equal to  $(r+1)^4 + 1$ . Let  $k \in \{0, \dots, L_n\}$ . We have :

$$\begin{aligned} \mathbb{E} \left[ \left( \sum_{\ell=\alpha+1}^{\alpha+\theta} \xi_\ell N_{n-1}(\ell) \right)^2 \middle| (S_p)_p \right] (\omega) &\leq \sum_{\ell=\alpha+1}^{\alpha+\theta} \sum_{m=\alpha+1}^{\alpha+\theta} |\mathbb{E}[\xi_\ell \xi_m]| N_{n-1}(\ell)(\omega) N_{n-1}(m)(\omega) \\ &\leq \theta \sum_{m \in \mathbb{Z}} |\mathbb{E}[\xi_0 \xi_m]| n^{1+2\delta_2}. \end{aligned}$$

Hence we have :

$$\begin{aligned}
\mathbb{E}[|b_k - 1| |(S_p)_p](\omega) &\leq n^{-\frac{1}{2} - \delta_3 + \delta_2} \left( \mathbb{E} \left[ \left| \sum_{y=\alpha(k)}^{\alpha(k+1)-1} \xi_y N_{n-1}(y) \right| \middle| (S_p)_p \right] (\omega) \right) \\
&\leq n^{-\frac{1}{2} - \delta_3 + \delta_2} n^{\frac{\beta}{2}} \sqrt{\sum_{m \in \mathbb{Z}} |\mathbb{E}[\xi_0 \xi_m]|} n^{\frac{1}{2} + \delta_2} \\
&\leq n^{-\delta_3 + 2\delta_2 + \frac{\beta}{2}} \sqrt{\sum_{m \in \mathbb{Z}} |\mathbb{E}[\xi_0 \xi_m]|} \\
&\leq n^{-\frac{3}{4}\delta_3 + \frac{3}{2}\delta_2} \sqrt{\sum_{m \in \mathbb{Z}} |\mathbb{E}[\xi_0 \xi_m]|}
\end{aligned}$$

since we have  $\beta < \frac{\delta_3}{2} - \delta_2$ . Moreover we have :

$$\begin{aligned}
|a_k(\omega) - 1| &\leq \frac{1}{2n^{1+2\delta_3-2\delta_2}} \sigma_\xi^2 \sum_{y=\alpha(k)}^{\alpha(k+1)} (N_{n-1}(y)(\omega))^2 \\
&\leq \frac{1}{2n^{1+2\delta_3-2\delta_2}} n^\beta \sigma_\xi^2 n^{1+2\delta_2} \\
&\leq \frac{1}{2} n^{-2\delta_3+4\delta_2+\beta} \sigma_\xi^2 \\
&\leq \frac{1}{2} n^{-\frac{3}{2}\delta_3+3\delta_2} \sigma_\xi^2.
\end{aligned}$$

From which, we get :

$$n^{\delta_2} \sum_{k=0}^{(r+1)^4-1} \mathbb{E}[|b_k - a_k| |(S_p)_p](\omega) + \mathbb{E}[|b_{L_n} - a_{L_n}| |(S_p)_p](\omega) \leq c_0 \left( n^{-\frac{3}{4}\delta_3 + \frac{5}{2}\delta_2} + n^{-\frac{3}{2}\delta_3 + 4\delta_2} \right), \quad (2)$$

with  $c_0 := ((r+1)^4 + 1) \sqrt{\sum_{m \in \mathbb{Z}} |\mathbb{E}[\xi_0 \xi_m]|} + \frac{1}{2} \sigma_\xi^2$ . Let us recall that  $\frac{5}{3}\delta_2 < \frac{1}{2}\delta_3$ .

Hence, it remains to estimate :

$$n^{\delta_2} \sum_{k=(r+1)^4}^{L_n-1} \left| \mathbb{E} \left[ \left( \prod_{m=0}^{k-1} b_m \right) (b_k - a_k) \prod_{m'=k+1}^{L_n} a_{m'} \middle| (S_p)_p \right] \right|. \quad (3)$$

- Let us control the following quantity :

$$\tilde{B}_n := n^{\delta_2} \sum_{k=(r+1)^4}^{L_n-1} \left| \mathbb{E} \left[ \left( \prod_{m=0}^{k-(r+1)^4} b_m \right) \prod_{j=1}^3 \left( \left( \prod_{m=k-(r+1)^{j+1}+1}^{k-(r+1)^j} b_m \right) - 1 \right) \prod_{m'=k-r}^{k-1} b_{m'} (b_k - a_k) \prod_{m'=k+1}^L a_{m'} \middle| (S_p)_p \right] \right|.$$

We have :

$$\tilde{B}_n(\omega) \leq n^{\delta_2} \sum_{k=(r+1)^4}^{L_n-1} \prod_{j=1}^3 \left\| \left( \prod_{m=k-(r+1)^{j+1}+1}^{k-(r+1)^j} b_m \right) - 1 \right\|_{L^\infty(U_n)} \|b_k - a_k\|_{L^\infty(U_n)}.$$

On  $U_n$ , we have :

$$|b_k - 1| \leq n^{-\frac{1}{2} - \delta_3 + \delta_2} n^\beta n^{\frac{1}{2} + \delta_2} \leq n^{-\delta_3 + 2\delta_2 + \beta}.$$

Analogously, we get :

$$\left| \left( \prod_{m=k-(r+1)^{j+1}+1}^{k-(r+1)^j} b_m \right) - 1 \right| \leq r(r+1)^j n^{-\delta_3 + 2\delta_2 + \beta}.$$



On the other hand, we have :

$$|a_k - 1| \leq \frac{1}{2} n^{-2\delta_3 + 4\delta_2 + \beta} \sigma_\xi^2.$$

Therefore we have :

$$\begin{aligned} \tilde{B}_n &\leq 4n^{\delta_2} n^{\frac{1}{2} + \delta_1 - \beta} r^3 (r+1)^6 \left(1 + \frac{1}{2} \sigma_\xi^2\right) (n^{-\delta_3 + 2\delta_2 + \beta})^4 \\ &\leq 4 \left(1 + \frac{1}{2} \sigma_\xi^2\right) r^3 (r+1)^6 n^{\frac{1}{2} - \frac{5}{2}\delta_3 + 6\delta_2 + \delta_1}, \end{aligned}$$

according to the fact that  $\beta < \frac{\delta_3}{2} - \delta_2$ . The control of the quantity  $\tilde{B}_n$  comes from the fact that  $\frac{5}{2}\delta_3 > \frac{1}{2} + 6\delta_2 + \delta_1$ .

- It remains to estimate :

$$n^{\delta_2} \sum_{k=(r+1)^4+1}^{L_n-1} \sum_{1 \leq j_0 < j_1 \leq j_2 \leq 4} C_{n,k,j_0,j_1,j_2}$$

where  $C_{n,k,j_0,j_1,j_2}$  is the following quantity :

$$\left| \mathbb{E} \left[ \left( \prod_{m=0}^{k-(r+1)^4} b_m \right) \left( \prod_{m=k-(r+1)^{j_2}+1}^{k-(r+1)^{j_1}} b_m \right) \left( \prod_{m=k-(r+1)^{j_0}+1}^{k-1} b_m \right) (b_k - a_k) \prod_{m'=k+1}^{L_n} a_{m'} \right] \right|,$$

with the convention :  $\prod_{m=\alpha}^{\beta} b_m = 1$  if  $\beta < \alpha$ .

- Let  $j_0, j_1, j_2$  be fixed. Let us control  $C_{n,k,j_0,j_1,j_2}$ . We have :

$$C_{n,k,j_0,j_1,j_2} \leq D_{n,k,j_0,j_1,j_2} + E_{n,k,j_0,j_1,j_2},$$

where :

$$D_{n,k,j_0,j_1,j_2} := \left| \text{Cov}_{|(S_p)_p} (\Delta_{n,k,j_1,j_2}, \Gamma_{n,k,j_0}) \prod_{m'=k+1}^{L_n} a_{m'} \right|,$$

and

$$E_{n,k,j_0,j_1,j_2} := \left| \mathbb{E}[\Delta_{n,k,j_1,j_2} | (S_p)_p] \mathbb{E}[\Gamma_{n,k,j_0} | (S_p)_p] \prod_{m'=k+1}^{L_n} a_{m'} \right|.$$

with  $\Delta_{n,k,j_1,j_2} := \prod_{m=0}^{k-(r+1)^4} b_m \prod_{m'=k-(r+1)^{j_2}+1}^{k-(r+1)^{j_1}} b_{m'}$  and  $\Gamma_{n,k,j_0} := \left( \prod_{m=k-(r+1)^{j_0}+1}^{k-1} b_m \right) (b_k - a_k)$ .

- Control of the covariance terms.

Let  $j_0, j_1, j_2$  be fixed. Let  $k = (r+1)^4, \dots, L_n - 1$ . We have :

$$\begin{aligned} D_{n,k,j_0,j_1,j_2} &\leq \left| \mathbb{E} \left[ \text{Cov}_{|(S_p)_p} \left( \Delta_{n,k,j_1,j_2}, \prod_{m=k-(r+1)^{j_0}+1}^k b_m \right) \prod_{m'=k+1}^{L_n} a_{m'} \right] \right| + \\ &\quad + \left| \mathbb{E} \left[ \text{Cov}_{|(S_p)_p} \left( \Delta_{n,k,j_1,j_2}, \prod_{m=k-(r+1)^{j_0}+1}^{k-1} b_m \right) \prod_{m'=k}^{L_n} a_{m'} \right] \right|. \end{aligned}$$

But we have :

$$\prod_{m=\theta_1+1}^{\theta_1+\theta_2} b_m = \exp \left( iun^{-\frac{1}{2}-\delta_3+\delta_2} \sum_{\ell=\alpha(\theta_1)}^{\alpha(\theta_1+\theta_2+1)-1} \xi_\ell N_{n-1}(\ell) \right).$$

Therefore, according to point 4 of the hypothesis of our theorem, we have :

$$D_{n,k,j_0,j_1,j_2} \leq 2C \left( 1 + n^{-\frac{1}{2}-\delta_3+\delta_2} \sum_{\ell \in \mathbb{Z}} N_{n-1}(\ell) \right) \varphi_{p,s}$$

with  $p := \lfloor n^\beta \rfloor ((r+1)^{j_1} - (r+1)^{j_0})$  and  $s := \lfloor n^\beta \rfloor (r+1)^{j_0} - 1$ . Let us notice that we have :  $p \geq rs$ . Hence we have :

$$\begin{aligned}
n^{\delta_2} \sum_{k=(r+1)^4}^{L_n-1} D_{n,k,j_0,j_1,j_2} &\leq 4C (n^{1-\delta_3+\delta_1-\beta+2\delta_2}) n^{-6\beta} \sup_{s \geq n^\beta} s^6 \varphi_{rs,s} \\
&\leq 4C \left( n^{1-\delta_3+\delta_1-7\frac{\delta_3}{2}+14\delta_2+2\delta_2} \right) \sup_{s \geq n^\beta} s^6 \varphi_{rs,s} \\
&\leq 4C \left( n^{1-\frac{\alpha}{2}\delta_3+\delta_1+16\delta_2} \right) \sup_{s \geq n^\beta} s^6 \varphi_{rs,s} \\
&\leq 4C \left( n^{1-\frac{\alpha}{8}+\delta_1+(\frac{27}{2}+16)\delta_2} \right) \sup_{s \geq n^\beta} s^6 \varphi_{rs,s},
\end{aligned}$$

since  $\beta > \frac{\delta_3}{2} - 2\delta_2$  and  $\delta_3 > \frac{1}{4} - 3\delta_2$ . We finish the control of these terms by noticing that  $\delta_1 + (\frac{27}{2} + 16)\delta_2 < \frac{1}{8}$ .

- Control of the term with the product of the expectations.

Let  $j_0, j_1, j_2$  be fixed. Let  $k = (r+1)^4, \dots, L_n - 1$ . We can notice that  $E_{n,k,j_0,j_1,j_2}$  is bounded from away by the following quantity :

$$F_{n,k,j_0} := \left| \mathbb{E} \left[ \prod_{m=k-(r+1)^{j_0}+1}^k b_m - \left( \prod_{m=k-(r+1)^{j_0}+1}^{k-1} b_m \right) a_k \middle| (S_p)_p \right] \right|.$$

We use Taylor expansions of the exponential function.

- First we explain that in  $F_{n,k,j_0}$ , we can replace

$$\prod_{m=k-(r+1)^{j_0}+1}^k b_m = \exp \left( iun^{-\frac{1}{2}-\delta_3+\delta_2} \sum_{\ell=\alpha_{(k-(r+1)^{j_0}+1)}}^{\alpha_{(k+1)}-1} \xi_\ell N_{n-1}(\ell) \right)$$

by the formula given by the Taylor expansion of the exponential function at the second order :

$$1 + iun^{-\frac{1}{2}-\delta_3+\delta_2} \sum_{\ell=\alpha_{(k-(r+1)^{j_0}+1)}}^{\alpha_{(k+1)}-1} \xi_\ell N_{n-1}(\ell) - \frac{u^2}{2n^{1+2\delta_3-2\delta_2}} \left( \sum_{\ell=\alpha_{(k-(r+1)^{j_0}+1)}}^{\alpha_{(k+1)}-1} \xi_\ell N_{n-1}(\ell) \right)^2. \quad (4)$$

Indeed, we control the error by :

$$\frac{1}{6} n^{-\frac{3}{2}-3\delta_3+3\delta_2} \mathbb{E} \left[ \left| \sum_{\ell=\alpha_{(k-(r+1)^{j_0}+1)}}^{\alpha_{(k+1)}-1} \xi_\ell N_{n-1}(\ell) \right|^3 \middle| (S_p)_p \right].$$

Moreover, we have :

$$\begin{aligned}
\mathbb{E} \left[ \left| \sum_{\ell=\alpha_{(k-(r+1)^{j_0}+1)}}^{\alpha_{(k+1)}-1} \xi_\ell N_{n-1}(\ell) \right|^4 \middle| (S_p)_p \right] &\leq \sum_{y_1, y_2, y_3, y_4 = \alpha_{(k-(r+1)^{j_0}+1)}}^{\alpha_{(k+1)}-1} \mathbb{E}[\xi_{y_1} \xi_{y_2} \xi_{y_3} \xi_{y_4}] \left( n^{\frac{1}{2}+\delta_2} \right)^4 \\
&\leq c'_0 n^{2+4\delta_2} (r+1)^6 n^{2\beta},
\end{aligned}$$

according to the hypothesis of our theorem. Hence, taking the sum over  $k = (r+1)^4, \dots, L_n - 1$  and multiplying by  $n^{\delta_2}$ , this substitution induces a total error bounded by :

$$\frac{1}{12} n^{\delta_2+\frac{1}{2}+\delta_1-\beta} n^{-\frac{3}{2}-3\delta_3+3\delta_2} n^{\frac{3}{2}+3\delta_2} (r+1)^{\frac{\alpha}{2}} n^{\frac{3}{2}\beta}$$

and so by :

$$\frac{1}{12} n^{7\delta_2 + \frac{1}{2} + \delta_1 - 3\delta_3 + \frac{1}{2}\beta} (r+1)^{\frac{9}{2}}.$$

But, using the facts that  $\beta < \frac{\delta_3}{2} - \delta_2$  and that  $\delta_3 > \frac{1}{4} - 3\delta_2$ , we have :

$$\begin{aligned} 7\delta_2 + \frac{1}{2} + \delta_1 - 3\delta_3 + \frac{1}{2}\beta &\leq \frac{13}{2}\delta_2 + \frac{1}{2} + \delta_1 - \frac{11}{4}\delta_3 \\ &\leq -\frac{3}{16} + \frac{59}{4}\delta_2 + \delta_1 \\ &\leq -\frac{1}{16}, \end{aligned}$$

since  $\delta_1 + (\frac{27}{2} + 16)\delta_2 < \frac{1}{8}$ .

– Let us introduce  $Y_k := \sum_{\ell=\alpha_{(k-(r+1)^{j_0+1})}^{\alpha(k)}-1}^{\alpha(k)-1} \xi_\ell N_{n-1}(\ell)$  and  $Z_k := \sum_{\ell=\alpha_{(k)}^{\alpha(k+1)}-1}^{\alpha(k+1)-1} \sigma_\xi^2 N_{n-1}(\ell)^2$ . We explain that, in  $F_{n,k,j_0}$ , we can replace

$$\left( \prod_{m=k-(r+1)^{j_0+1}}^{k-1} b_m \right) a_k = \exp \left( \frac{i u}{n^{\frac{1}{2} + \delta_3 - \delta_2}} Y_k - \frac{u^2}{2n^{1+2\delta_3-2\delta_2}} Z_k \right)$$

by the formula given by the Taylor expansion of the exponential function at the second order :

$$1 + \frac{i u}{n^{\frac{1}{2} + \delta_3 - \delta_2}} Y_k - \frac{u^2}{2n^{1+2\delta_3-2\delta_2}} Z_k + \frac{1}{2} \left( \frac{i u}{n^{\frac{1}{2} + \delta_3 - \delta_2}} Y_k - \frac{u^2}{2n^{1+2\delta_3-2\delta_2}} Z_k \right)^2. \quad (5)$$

Indeed the modulus of the error between these two quantities is less than :

$$\frac{1}{6} \mathbb{E} \left[ \left| \frac{i u}{n^{\frac{1}{2} + \delta_3 - \delta_2}} Y_k - \frac{u^2}{2n^{1+2\delta_3-2\delta_2}} Z_k \right|^3 \middle| (S_p)_p \right] \leq \frac{4}{3} \mathbb{E} \left[ \left| \frac{1}{n^{\frac{1}{2} + \delta_3 - \delta_2}} Y_k \right|^3 + \left| \frac{1}{2n^{1+2\delta_3-2\delta_2}} Z_k \right|^3 \middle| (S_p)_p \right].$$

The first term will be controled as just before. Let us control the second term. We have :

$$\begin{aligned} |n^{-1-2\delta_3+2\delta_2} Z_k|^3 &\leq n^{-3-6\delta_3+6\delta_2} (\sigma_\xi^2)^3 n^{3\beta} n^{3+6\delta_2} \\ &\leq n^{-6\delta_3+12\delta_2+3\beta} (\sigma_\xi^2)^3. \end{aligned}$$

Hence, taking the sum over  $k = (r+1)^4, \dots, L_n - 1$  and multiplying by  $n^{\delta_2}$ , we get a quantity bounded by :

$$2n^{\frac{1}{2} + \delta_1 - 6\delta_3 + 13\delta_2 + 2\beta} (\sigma_\xi^2)^3.$$

But we have :

$$\begin{aligned} \frac{1}{2} + \delta_1 - 6\delta_3 + 13\delta_2 + 2\beta &\leq \frac{1}{2} + \delta_1 - 5\delta_3 + 11\delta_2 \\ &\leq \frac{1}{2} - \frac{5}{4} + \delta_1 + 23\delta_2 < 0. \end{aligned}$$

since  $\beta < \delta_3 - 2\delta_2$  and  $\delta_3 > \frac{1}{4} - 3\delta_2$ .

– Now, we show that in formula (5), we can ommit the term with  $(Z_k)^2$ . Indeed, we have :

$$\begin{aligned} n^{\delta_2} \sum_{(r+1)^4}^{L_n-1} (n^{-1-2\delta_3+2\delta_2} Z_k)^2 &\leq 2n^{\delta_2 + \frac{1}{2} + \delta_1 - \beta - 2 - 4\delta_3 + 4\delta_2} n^{2\beta} (\sigma_\xi^2)^2 n^{2+4\delta_2} \\ &\leq 2n^{\frac{1}{2} + \delta_1 - 4\delta_3 + 9\delta_2 + \beta} (\sigma_\xi^2)^2 \\ &\leq 2n^{\frac{1}{2} + \delta_1 - \frac{7}{2}\delta_3 + 8\delta_2} (\sigma_\xi^2)^2 \\ &\leq 2n^{-\frac{1}{12} - \frac{1}{6}\delta_1 - \frac{1}{6}\delta_2} (\sigma_\xi^2)^2 \end{aligned}$$

since  $\beta < \frac{\delta_3}{2} - \delta_2$  and  $3\delta_3 > \frac{1}{2} + 7\delta_2 + \delta_1$ .

– Hence, it remains to estimate the following quantity called  $G_{n,k,j_0}$  :

$$\left| \mathbb{E} \left[ \frac{i u}{n^{\frac{1}{2} + \delta_3 - \delta_2}} (Y_k + W_k) - \frac{u^2}{2n^{1+2\delta_3-2\delta_2}} (Y_k + W_k)^2 - \frac{i u}{n^{\frac{1}{2} + \delta_3 - \delta_2}} Y_k + \frac{u^2}{2n^{1+2\delta_3-2\delta_2}} Z_k + \frac{u^2}{2n^{1+2\delta_3-2\delta_2}} (Y_k)^2 + \frac{i u}{n^{\frac{1}{2} + \delta_3 - \delta_2}} Y_k \frac{u^2}{2n^{1+2\delta_3-2\delta_2}} Z_k \middle| (S_p)_p \right] \right|$$

with  $W_k := \sum_{\ell=\alpha(k)}^{\alpha(k+1)-1} \xi_\ell N_{n-1}(\ell)$ . We get :

$$\begin{aligned} G_{n,k,j_0} &= \left| \mathbb{E} \left[ -\frac{u^2}{2n^{1+2\delta_3-2\delta_2}} (Y_k + W_k)^2 + \frac{u^2}{2n^{1+2\delta_3-2\delta_2}} Z_k + \frac{u^2}{2n^{1+2\delta_3-2\delta_2}} (Y_k)^2 \middle| (S_p)_p \right] \right| \\ &= \frac{u^2}{2n^{1+2\delta_3-2\delta_2}} \left| \mathbb{E} \left[ (W_k)^2 + 2W_k Y_k - Z_k \middle| (S_p)_p \right] \right|. \end{aligned}$$

Let us notice that we have :

$$Z_k := \sum_{\ell=\alpha(k)}^{\alpha(k+1)-1} \left( \mathbb{E}[(\xi_\ell)^2] N_{n-1}(\ell)^2 + 2 \sum_{m \leq \ell-1} \mathbb{E}[\xi_\ell \xi_m] N_{n-1}(\ell)^2 \right).$$

– Let us show that, in the last expression of  $G_{n,k,j_0}$ , we can replace  $Z_k$  by :

$$\tilde{Z}_k := \sum_{\ell=\alpha(k)}^{\alpha(k+1)-1} \left( \mathbb{E}[(\xi_\ell)^2] N_{n-1}(\ell)^2 + 2 \sum_{m \leq \ell-1} \mathbb{E}[\xi_\ell \xi_m] N_{n-1}(\ell) N_{n-1}(m) \right).$$

Indeed, we have :

$$\begin{aligned} \frac{u^2}{2n^{1+2\delta_3-2\delta_2}} \mathbb{E} \left[ \left| Z_k - \tilde{Z}_k \right| \middle| (S_p)_p \right] &\leq \frac{1}{n^{1+2\delta_3-2\delta_2}} \sum_{\ell=\alpha(k)}^{\alpha(k+1)-1} \sum_{m \leq \ell-1} |\mathbb{E}[\xi_\ell \xi_m]| N_{n-1}(\ell) |N_{n-1}(m) - N_{n-1}(\ell)| \\ &\leq \frac{1}{n^{1+2\delta_3-2\delta_2}} n^\beta \sum_{m \geq 1} |\mathbb{E}[\xi_0 \xi_m]| n^{\frac{1}{2} + \delta_2} \sqrt{m} n^{\frac{1}{4} + \frac{\gamma}{2}} \\ &\leq n^{-\frac{1}{4} - 2\delta_3 + 3\delta_2 + \beta + \frac{\gamma}{2}} \sum_{m \geq 1} \sqrt{m} |\mathbb{E}[\xi_0 \xi_m]|. \end{aligned}$$

Hence, taking the sum over  $k = (r+1)^4, \dots, L_n - 1$  and multiplying by  $n^{\delta_2}$ , we get a quantity bounded by :

$$4n^{\frac{1}{4} + \delta_1 - 2\delta_3 + 4\delta_2 + \frac{\gamma}{2}} \sum_{m \geq 1} \sqrt{m} |\mathbb{E}[\xi_0 \xi_m]|.$$

But we have :

$$\frac{1}{4} + \delta_1 - 2\delta_3 + 4\delta_2 + \frac{\gamma}{2} \leq -\frac{1}{4} + \delta_1 + 10\delta_2 + \frac{\gamma}{2} < 0$$

since  $\delta_3 > \frac{1}{4} - 3\delta_2$  and  $\gamma < \frac{1}{2} - 22 \max(\delta_1, \delta_2)$ .

– Hence we have to estimate :

$$\tilde{G}_{n,k,j_0} = \frac{u^2}{2n^{1+2\delta_3-2\delta_2}} \left| \mathbb{E} \left[ (W_k)^2 + 2W_k Y_k - \tilde{Z}_k \middle| (S_p)_p \right] \right|.$$

We have :

$$\mathbb{E} \left[ (W_k)^2 \middle| (S_p)_p \right] = \sum_{\ell=\alpha(k)}^{\alpha(k+1)-1} \left( \mathbb{E}[(\xi_\ell)^2] (N_{n-1}(\ell))^2 + 2 \sum_{m=\alpha(k)}^{\ell-1} \mathbb{E}[\xi_\ell \xi_m] N_{n-1}(\ell) N_{n-1}(m) \right).$$

Hence we have :

$$\mathbb{E} \left[ (W_k)^2 + 2W_k Y_k \middle| (S_p)_p \right] = \sum_{\ell=\alpha(k)}^{\alpha(k+1)-1} \left( \mathbb{E}[(\xi_\ell)^2] (N_{n-1}(\ell))^2 + 2 \sum_{m=\alpha_{k-(r+1)^{j_0}+1}}^{\ell-1} \mathbb{E}[\xi_\ell \xi_m] N_{n-1}(\ell) N_{n-1}(m) \right).$$

We get :

$$\begin{aligned}
\tilde{G}_{n,k,j_0} &= \frac{u^2}{2n^{1+2\delta_3-2\delta_2}} \left| \sum_{\ell=\alpha(k)}^{\alpha(k+1)-1} \sum_{m \leq \alpha_{k-(r+1)j_0+1}^{-1}} \mathbb{E}[\xi_\ell \xi_m] N_{n-1}(\ell) N_{n-1}(m) \right| \\
&\leq \frac{u^2}{2n^{1+2\delta_3-2\delta_2}} n^\beta \sum_{m \geq \frac{(r+1)n^\beta}{2}} |\mathbb{E}[\xi_0 \xi_m]| n^{1+2\delta_2} \\
&\leq \frac{1}{2} n^{-2\delta_3+4\delta_2+\beta} \sum_{m \geq \frac{(r+1)n^\beta}{2}} |\mathbb{E}[\xi_0 \xi_m]|.
\end{aligned}$$

Hence, taking the sum over  $k = (r+1)^4, \dots, L_n - 1$  of these quantities and multiplying by  $n^{\delta_2}$ , we get a quantity bounded by :

$$2n^{\frac{1}{2}+\delta_1-2\delta_3+5\delta_2} \sum_{m \geq \frac{(r+1)n^\beta}{2}} |\mathbb{E}[\xi_0 \xi_m]| \leq 2n^{\delta_1+11\delta_2} \sum_{m \geq \frac{(r+1)n^\beta}{2}} |\mathbb{E}[\xi_0 \xi_m]|,$$

since  $\delta_3 > \frac{1}{4} - 3\delta_2$ . To conclude it suffices to notice that :

$$n^{\delta_1+11\delta_2} \sum_{m \geq \frac{(r+1)n^\beta}{2}} |\mathbb{E}[\xi_0 \xi_m]| = O(n^{-\varepsilon}).$$

### 3.1.2 Proof of lemma 5

Let us consider  $n \geq 2$ . According to lemma 6, it suffices to prove that there exists a real number  $\delta' > 0$  such that we have :

$$\sup_{n \geq 1} n^{\delta'} \sup_{\omega \in U_n} n^{\delta_2} \int_{|u| \leq 1} \exp\left(-\frac{u^2}{2n^{1+2\delta_3-2\delta_2}} \sum_{y,z} \mathbb{E}[\xi_y \xi_z] (N_{n-1}(y)(\omega))^2\right) e^{-\frac{u^2 n^{2\delta_2}}{2}} du < +\infty.$$

Let us take  $\omega \in U_n$ . We have :

$$\exp\left(-\frac{u^2}{2n^{1+2\delta_3-2\delta_2}} \sum_{y,z} \mathbb{E}[\xi_y \xi_z] (N_{n-1}(y)(\omega))^2\right) = \exp\left(-\frac{u^2}{2n^{1+2\delta_3-2\delta_2}} \sigma_\xi^2 \sum_y (N_{n-1}(y)(\omega))^2\right).$$

Let us define :  $p_n := \text{Card}\{y \in \mathbb{Z} : N_{n-1}(y) \geq \frac{n^{\frac{1}{2}-\delta_4}}{3}\}$ . We have :

$$\begin{aligned}
n &= \sum_{y=-\lfloor n^{\frac{1}{2}+\delta_1} \rfloor}^{\lfloor n^{\frac{1}{2}+\delta_1} \rfloor} N_{n-1}(y) \\
&\leq p_n n^{\frac{1}{2}+\delta_2} + \frac{n^{\frac{1}{2}-\delta_4}}{3} (3n^{\frac{1}{2}+\delta_1} - p_n) \\
&\leq p_n \left( n^{\frac{1}{2}+\delta_2} - \frac{n^{\frac{1}{2}-\delta_4}}{3} \right) + n^{\frac{1}{2}-\delta_4} n^{\frac{1}{2}+\delta_1} \\
&\leq p_n n^{\frac{1}{2}+\delta_2} \left( 1 - \frac{n^{-(\delta_2+\delta_4)}}{3} \right) + n^{1+\delta_1-\delta_4}.
\end{aligned}$$

Let us recall that :  $\delta_1 < \delta_4 < \frac{1}{4} - \delta_3 - \frac{5}{2}\delta_2$ . Hence we have :

$$p_n \geq n^{-\frac{1}{2}-\delta_2} (n - n^{1-(\delta_4-\delta_1)}) \geq n^{\frac{1}{2}-\delta_2} (1 - n^{-(\delta_4-\delta_1)}) \geq c_0 n^{\frac{1}{2}-\delta_2},$$

with  $c_0 := 1 - 2^{-(\delta_4 - \delta_1)}$ .

Hence we have :

$$\begin{aligned} \sum_{y \in \mathbb{Z}} (N_{n-1}(y)(\omega))^2 &\geq p_n \left( \frac{n^{\frac{1}{2} - \delta_4}}{3} \right)^2 \geq \frac{c_0 n^{\frac{3}{2} - \delta_2 - 2\delta_4}}{9}. \\ \exp \left( -\frac{u^2}{2n^{1+2\delta_3-2\delta_2}} \sum_y \sigma_\xi^2 (N_{n-1}(y)(\omega))^2 \right) &\leq \exp \left( -\frac{u^2}{18n^{1+2\delta_3-2\delta_2}} \sigma_\xi^2 c_0 n^{\frac{3}{2} - \delta_2 - 2\delta_4} \right) \\ &\leq \exp \left( -\frac{u^2}{18} \sigma_\xi^2 c_0 n^{\frac{1}{2} - 3\delta_2 - 2\delta_3 - 2\delta_4} \right). \end{aligned}$$

Therefore, we have :

$$\begin{aligned} n^{\delta_2} \int_{|u| \leq 1} \exp \left( -\frac{u^2}{2n^{1+2\delta_3-2\delta_2}} \sum_{y,z} \mathbb{E}[\xi_y \xi_z] (N_{n-1}(y)(\omega))^2 \right) e^{-\frac{u^2 n^{2\delta_2}}{2}} du &\leq \\ &\leq n^{\delta_2} \int_{|u| \leq 1} \exp \left( -\frac{u^2}{18} \sigma_\xi^2 c_0 n^{\frac{1}{2} - 3\delta_2 - 2\delta_3 - 2\delta_4} \right) du \\ &\leq \frac{n^{\delta_2}}{n^{\frac{1}{4} - \delta_4 - \frac{3}{2}\delta_2 - \delta_3}} \int_{\mathbb{R}} \exp \left( -\frac{v^2}{18} \sigma_\xi^2 c_0 \right) dv \\ &\leq n^{-\frac{1}{4} + \delta_4 + \frac{5}{2}\delta_2 + \delta_3} \int_{\mathbb{R}} \exp \left( -\frac{v^2}{18} \sigma_\xi^2 c_0 \right) dv. \end{aligned}$$

with the change of variable  $v = un^{\frac{1}{4} - \delta_4 - \frac{3}{2}\delta_2 - \delta_3}$ . This gives the result since  $\delta_4 + \delta_3 + \frac{5}{2}\delta_2 < \frac{1}{4}$  *qed*.

## 4 About the model of Guillotin-Le Ny

In this section, we prove that the hypothesis  $\int_M \frac{1}{\sqrt{f(1-f)}} d\nu < +\infty$  of Guillotin-Le Ny in [2] can be replaced by  $\int_M \frac{1}{(f(1-f))^p} d\nu < +\infty$ , for some  $p > 0$ . In this situation, there is no need to introduce the set  $U_n$ ; we can take  $U_n = A_n$ . If we take  $\delta_1 > 0$ ,  $\delta_2 > 0$  and  $\delta_3 > 0$ , all the points except the point 3(b)(ii) comes in the same way without the need of the hypothesis  $\int_M \frac{1}{\sqrt{f(1-f)}} d\nu < +\infty$ .

Guillotin and Le Ny have to estimate the following quantity :

$$\sup_{\omega \in A_n} n^{\frac{1}{2} + \delta_3} I_n^{(1)}(\omega) := \int_{\{|t| \leq n^{-\frac{1}{2} - \delta_3 + \delta_2}\}} \mathbb{E} \left[ e^{it \sum_{y \in \mathbb{Z}} \xi_y N_{n-1}(y)} \middle| (S_p)_p \right] (\omega) e^{-\frac{t^2 n^{1+2\delta_3}}{2}} dt.$$

Let us take  $\omega \in A_n$ .

We will suppose  $\delta_3 > 2\delta_2$  and  $\delta_1 < \delta_4 < \frac{1}{4} - \delta_3 - \frac{\delta_2}{2}$ .

The idea of Guillotin and Le Ny is to write :

$$\begin{aligned} n^{\frac{1}{2} + \delta_3} |I_n^{(1)}| &\leq n^{\frac{1}{2} + \delta_3} \int_{\{|t| \leq n^{-\frac{1}{2} - \delta_3 + \delta_2}\}} \mathbb{E} \left[ \prod_{y \in \mathbb{Z}} |\cos(tN_{n-1}(y)) + i(2f \circ T^y - 1) \sin(tN_{n-1}(y))| \middle| (S_p)_p \right] e^{-\frac{t^2 n^{1+2\delta_3}}{2}} dt \\ &\leq n^{\frac{1}{2} + \delta_3} \int_{\{|t| \leq n^{-\frac{1}{2} - \delta_3 + \delta_2}\}} \mathbb{E} \left[ \prod_{y \in \mathbb{Z}} \sqrt{1 - 4f \circ T^y (1 - f \circ T^y) \sin^2(tN_{n-1}(y))} \middle| (S_p)_p \right] e^{-\frac{t^2 n^{1+2\delta_3}}{2}} dt \\ &\leq n^{\frac{1}{2} + \delta_3} \int_{\{|t| \leq n^{-\frac{1}{2} - \delta_3 + \delta_2}\}} \mathbb{E} \left[ \prod_{y \in \mathbb{Z}} \sqrt{1 - 4f \circ T^y (1 - f \circ T^y) \frac{16}{\pi^2} (tN_{n-1}(y))^2} \middle| (S_p)_p \right] e^{-\frac{t^2 n^{1+2\delta_3}}{2}} dt \end{aligned}$$

$$\leq n^{\frac{1}{2}+\delta_3} \int_{\{|t| \leq n^{-\frac{1}{2}-\delta_3+\delta_2}\}} \mathbb{E} \left[ \prod_{y \in \mathbb{Z}} e^{-8\pi^2 f \circ T^y (1-f \circ T^y) t^2 N_{n-1}(y)^2} |(S_p)_p \right] e^{-\frac{t^2 n^{1+2\delta_3}}{2}} dt$$

since  $|tN_{n-1}(y)| \leq n^{-\frac{1}{2}-\delta_3+\delta_2} n^{\frac{1}{2}+\delta_2} = n^{2\delta_2-\delta_3}$ . Hence, if  $n$  is large enough, then  $|tN_{n-1}(y)|$  will be uniformly less than  $\frac{\pi}{2}$  and  $|\sin(tN_{n-1}(y))| \geq \frac{2}{\pi}|tN_{n-1}(y)|$ . We also use the fact that, for positive  $u$ , we have :  $1 - u \leq e^{-u}$ . According to the Hölder inequality with  $\sum_y \frac{N_{n-1}(y)^2}{\sum_k N_{n-1}(k)^2} = 1$ , we have :

$$n^{\frac{1}{2}+\delta_3} \left| I_n^{(1)} \right| \leq n^{\frac{1}{2}+\delta_3} \int_{\mathbb{R}} \mathbb{E} \left[ e^{-8\pi^2 f(1-f)t^2 \sum_k N_{n-1}(k)^2} |(S_p)_p \right] e^{-\frac{t^2 n^{1+2\delta_3}}{2}} dt.$$

Now, we use the fact that, since  $\delta_4 > \delta_1$ , there exists a constant  $c$  such that we have :

$$\forall \omega' \in A_n, \quad \sum_y (N_{n-1}(y))^2(\omega') \geq cn^{\frac{3}{2}-\delta_2-2\delta_4}.$$

This has been proved in the previous section called 'proof of lemma 5'. Hence, under the hypothesis  $\int_M \frac{1}{\sqrt{f(1-f)}} d\nu < +\infty$  of Guillotin and Le Ny, we have :

$$\begin{aligned} n^{\frac{1}{2}+\delta_3} \left| I_n^{(1)}(\omega) \right| &\leq n^{\frac{1}{2}+\delta_3} \int_{\{|t| \leq n^{-\frac{1}{2}-\delta_3+\delta_2}\}} \mathbb{E} \left[ e^{-8\pi^2 f(1-f)t^2 n^{\frac{3}{2}-\delta_2-2\delta_4}} \right] e^{-\frac{t^2 n^{1+2\delta_3}}{2}} dt \\ &\leq n^{-\frac{1}{4}+\delta_3+\frac{\delta_2}{2}+\delta_4} \int_{\mathbb{R}} \mathbb{E} \left[ \frac{1}{\sqrt{f(1-f)}} \right] e^{-8\pi^2 v^2} dv \end{aligned}$$

with the change of variable  $v = t\sqrt{f(1-f)n^{\frac{3}{2}-\delta_2-2\delta_4}}$ . We adapt this argument to our hypothesis. Now let us replace the hypothesis  $\int_M \frac{1}{\sqrt{f(1-f)}} d\nu < +\infty$  by  $\int_M \frac{1}{[f(1-f)]^p} d\nu < +\infty$  for some  $p > 0$ . Let us take  $\delta_1 < \delta_4 < \frac{1}{4} - \delta_3 - \frac{\delta_2}{2} - \frac{\delta_2}{p}$ . We use the fact that there exists a constant  $c_p > 0$  such that, for any real number  $u > 0$ , we have :  $e^{-u} \leq \frac{c_p}{u^p}$ . We have :

$$\begin{aligned} n^{\frac{1}{2}+\delta_3} \int_{\{|t| \leq n^{-\frac{1}{4}+\frac{\delta_2}{2}+\delta_4+\frac{\delta_2}{p}}\}} \mathbb{E} \left[ e^{-8\pi^2 f(1-f)t^2 n^{\frac{3}{2}-\delta_2-2\delta_4}} \right] e^{-\frac{t^2 n^{1+2\delta_3}}{2}} dt &\leq n^{\frac{1}{2}+\delta_3} n^{-\frac{3}{4}+\frac{\delta_2}{2}+\delta_4+\frac{\delta_2}{p}} \\ &\leq n^{-\frac{1}{4}+\delta_3+\frac{\delta_2}{2}+\delta_4+\frac{\delta_2}{p}}. \end{aligned}$$

On the other hand, we have :

$$\begin{aligned} n^{\frac{1}{2}+\delta_3} \int_{\{n^{-\frac{3}{4}+\frac{\delta_2}{2}+\delta_4+\frac{\delta_2}{p}} < |t| < n^{-\frac{1}{2}-\delta_3+\delta_2}\}} \mathbb{E} \left[ e^{-8\pi^2 f(1-f)t^2 n^{\frac{3}{2}-\delta_2-2\delta_4}} \right] e^{-\frac{t^2 n^{1+2\delta_3}}{2}} dt &\leq \\ &\leq n^{\frac{1}{2}+\delta_3} n^{-\frac{1}{2}-\delta_3+\delta_2} \int_M e^{-8\pi^2 f(1-f)n^{\frac{2\delta_2}{p}}} d\nu \\ &\leq n^{-\delta_2} c_p \left( \frac{1}{8\pi^2} \right)^p \int_M [f(1-f)]^{-p} d\nu. \end{aligned}$$

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